

Homework #4 - Solutions

Problem 1. (a) If $f: \mathbb{R}^n \rightarrow \mathbb{R}^n$ is one-to-one, onto and differentiable, then the change of variables formula of multivariate calculus states that, for any integrable function φ on \mathbb{R}^n ,

$$\int_{\mathbb{R}^n} \varphi(y) d^n y = \int_{\mathbb{R}^n} \varphi(f(x)) \left| \det \left(\frac{\partial f}{\partial x}(x) \right) \right| d^n x.$$

We apply this result to the function

$$\varphi(y) = \tilde{1}_B(y) P_Y(y)$$

to obtain

$$\int_B P_Y(y) d^n y = \int_{\mathbb{R}^n} \tilde{1}_B(f(x)) P_Y(f(x)) \left| \det \left(\frac{\partial f}{\partial x}(x) \right) \right| d^n x.$$

By definition of the density $P_Y(y)$,

$$\begin{aligned} \int_B P_Y(y) d^n y &= P(\tilde{Y} \in B) \\ &= P(f(\tilde{X}) \in B) \\ &= P(\tilde{X} \in f^{-1}(B)) = P(\tilde{X} \in A) \end{aligned}$$

Also,

$$\tilde{1}_B(f(x)) = \tilde{1}_A(x)$$

since

$$f(x) \in B \iff x \in f^{-1}(B) = A.$$

We thus see that

$$P(\tilde{X} \in A) = \int_A P_Y(f(x)) \left| \det \left(\frac{\partial f}{\partial x}(x) \right) \right| d^4x.$$

It therefore follows by the definition of the density that

$$P_X(x) = P_Y(f(x)) \left| \det \left(\frac{\partial f}{\partial x}(x) \right) \right|.$$

Substituting $y = f(x)$, this is equivalent to

$$P_X(f^{-1}(y)) = P_Y(y) \left| \det \left(\frac{\partial f}{\partial x}(f^{-1}(y)) \right) \right|$$

or,

$$P_Y(y) = \frac{P_X(f^{-1}(y))}{\left| \det \left(\frac{\partial f}{\partial x}(f^{-1}(y)) \right) \right|}.$$

(b) We apply the previous result with

$$x = \begin{pmatrix} x \\ y \end{pmatrix} = \begin{pmatrix} u+v \\ \frac{1}{2}(u-v) \end{pmatrix} = f(u)$$

for $u = \begin{pmatrix} u \\ v \end{pmatrix}$. Note that

$$\frac{\partial f}{\partial u} = \begin{pmatrix} \partial x / \partial u & \partial x / \partial v \\ \partial y / \partial u & \partial y / \partial v \end{pmatrix}$$

$$= \begin{pmatrix} 1 & 1 \\ \frac{1}{2} & -\frac{1}{2} \end{pmatrix}$$

So that

$$\det\left(\frac{\partial f}{\partial u}\right) = (1)\left(-\frac{1}{2}\right) - (1)\left(\frac{1}{2}\right) = -1.$$

Furthermore, solving the equations

$$u + v = x$$

$$u - v = 2y$$

for u & v gives

$$u = \frac{1}{2}x + y, \quad v = \frac{1}{2}x - y$$

or

$$u = \begin{pmatrix} u \\ v \end{pmatrix} = \begin{pmatrix} \frac{1}{2}x + y \\ \frac{1}{2}x - y \end{pmatrix} = f^{-1}(x).$$

We thus see that

$$P_{\mathbf{x}}(x) = P_{\mathbf{u}}(f^{-1}(x)) / \left| \det\left(\frac{\partial f}{\partial u}(f^{-1}(x))\right) \right|$$

or that

$$P_{X,Y}(x,y) = P_{U,V}(u,v) / |-1|$$

$$= P_{U,V}\left(\frac{1}{2}x+y, \frac{1}{2}x-y\right) \checkmark$$

Since

$$P_X(x) = \int P_{X,Y}(x,y) dy$$

it follows that

$$P_X(x) = \int P_{U,V}\left(\frac{1}{2}x+y, \frac{1}{2}x-y\right) dy$$

$$= \int P_{U,V}(u, x-u) du \checkmark$$

using $u = \frac{1}{2}x+y$ so that $du = dy$ (at fixed x) and

$$x-u = x - \left(\frac{1}{2}x+y\right) = \frac{1}{2}x-y.$$

Remark: If \tilde{U}, \tilde{V} are independent random variables, notice that the sum variable $\tilde{X} = \tilde{U} + \tilde{V}$ has the PDF

$$P_{\tilde{X}}(x) = \int P_{\tilde{U}}(u) P_{\tilde{V}}(x-u) du$$

$$= (P_{\tilde{U}} * P_{\tilde{V}})(x)$$

where " $*$ " denotes the convolution product of functions.

Problem 2. (a) Using the definition of the gamma function

$$\Gamma(z) = \int_0^{\infty} t^{z-1} e^{-t} dt,$$

we see by taking $z = a$ and $t = bx$, that

$$\Gamma(a) = b^a \int_0^{\infty} x^{a-1} e^{-bx} dx$$

or that

$$\int_0^{\infty} x^{a-1} e^{-bx} dx = \frac{\Gamma(a)}{b^a}.$$

We now take p derivatives of both sides of this equation. First,

$$\begin{aligned} \frac{\partial^p}{\partial b^p} \left(\int_0^{\infty} x^{a-1} e^{-bx} dx \right) &= \int_0^{\infty} x^{a-1} (-x)^p e^{-bx} dx \\ &= (-1)^p E(\tilde{X}^p) \cdot \frac{\Gamma(a)}{b^p} \end{aligned}$$

On the other hand,

$$\begin{aligned} \frac{\partial^p}{\partial b^p} \left(\frac{\Gamma(a)}{b^a} \right) &= (-a)(-a-1)\dots(-a-p+1) \frac{\Gamma(a)}{b^{a+p}} \\ &= (-1)^p \frac{(a)^p}{b^p} \cdot \frac{\Gamma(a)}{b^p} \end{aligned}$$

so that

$$E(\tilde{X}^p) = (a)^p / b^p. \quad \checkmark$$

Another approach is to note that

$$\begin{aligned} E(\tilde{X}^p) &= \frac{b^a}{\Gamma(a)} \int_0^{\infty} x^p \cdot x^{a-1} e^{-bx} dx \\ &= \frac{b^a}{\Gamma(a)} \int_0^{\infty} x^{a+p-1} e^{-bx} dx \\ &= \frac{b^a}{\Gamma(a)} \cdot \frac{\Gamma(a+p)}{b^{a+p}} = \frac{\Gamma(a+p)/\Gamma(a)}{b^p} . \end{aligned}$$

A basic property of the gamma function is that

$$\Gamma(z+1) = z\Gamma(z) .$$

Applying this p times gives

$$\Gamma(a+p) = (a+p-1)(a+p-2)\cdots a \Gamma(a)$$

or

$$\begin{aligned} \frac{\Gamma(a+p)}{\Gamma(a)} &= a(a+1)\cdots(a+p-1) \\ &= (a)^p , \end{aligned}$$

so that

$$E(\tilde{X}^p) = \frac{(a)^p}{b^p} . \quad \checkmark$$

b) We use the change of variables formula with

$$x = g(y) = 1/y$$

where

$$P_Y(y) = y^{a-1} \frac{b^a}{\Gamma(a)} e^{-by}, \quad y > 0,$$

Since

$$\frac{dx}{dy} = -\frac{1}{y^2},$$

we get from the change of variables formula that

$$\begin{aligned} P_X(x) &= \frac{P_Y(y)}{|dx/dy|} = y^2 P_Y(y) \Big|_{y=1/x} \\ &= \frac{1}{x^2} \cdot \left(\frac{1}{x}\right)^{a-1} \frac{b^a}{\Gamma(a)} e^{-b/x} \\ &= \frac{b^a}{\Gamma(a)} x^{-(a+1)} e^{-b/x}. \quad \checkmark \end{aligned}$$

c) Taking $a = \frac{1}{2}$, $b = \frac{1}{2}$ in the gamma density gives

$$p(x) = P_{1/2, 1/2}(x) = x^{\frac{1}{2}-1} \frac{(1/2)^{1/2}}{\Gamma(1/2)} e^{-x/2}.$$

We must furthermore use the result that

$$\Gamma\left(\frac{1}{2}\right) = \sqrt{\pi},$$

so that

$$p(x) = \frac{1}{\sqrt{2\pi x}} e^{-x/2}, \quad x > 0 \quad \checkmark$$

We next show that $X = \tilde{N}^2$ with $\tilde{N} \sim N(0, 1)$ is chi-squared.
 Taking $x = n^2$, we see that the inverse is double-valued:

$$n_{\pm} = \pm \sqrt{x},$$

with

$$\frac{dn_{\pm}}{dx} = \pm \frac{1}{2} x^{-1/2}.$$

Thus, we must use the generalized change of variables formula

$$p_X(x) = P_N(n_+) \left| \frac{dn_+}{dx} \right| + P_N(n_-) \left| \frac{dn_-}{dx} \right|$$

with

$$P_N(n) = \frac{1}{\sqrt{2\pi}} e^{-n^2/2}.$$

This gives

$$\begin{aligned} p_X(x) &= \frac{1}{\sqrt{2\pi}} e^{-x/2} \cdot \frac{1}{2} x^{-1/2} + \frac{1}{\sqrt{2\pi}} e^{-x/2} \cdot \frac{1}{2} x^{-1/2} \\ &= \frac{1}{\sqrt{2\pi}} x^{-1/2} e^{-x/2}, \quad \checkmark \end{aligned}$$

which is the chi-squared PDF.

Since $\tilde{X} > 0$ with probability one,

$$F_{\tilde{X}}(x) = P(X < x) = 0$$

for any $x < 0$. For $x > 0$ we get that

$$F_{\tilde{X}}(x) = \int_0^x du \frac{1}{\sqrt{2\pi u}} e^{-u/2}.$$

Then set

$$u = 2t^2$$

$$\sqrt{2u} = 2t$$

$$du = 4t dt$$

giving

$$F_{\tilde{X}}(x) = \int_0^{\sqrt{x/2}} 4t dt \cdot \frac{1}{2t\sqrt{\pi}} e^{-t^2}$$

$$= \frac{2}{\sqrt{\pi}} \int_0^{\sqrt{x/2}} dt e^{-t^2} = \text{erf}\left(\sqrt{\frac{x}{2}}\right) \quad \checkmark$$

(d) Setting $a = \frac{1}{2}$, $b = \frac{1}{2}$ in the inverse gamma PDF gives

$$P_{1/2, 1/2}(x) = \frac{(\frac{1}{2})^{1/2}}{\Gamma(\frac{1}{2})} \cdot x^{-(1+\frac{1}{2})} e^{-1/2x}$$

$$= \frac{1}{\sqrt{2\pi}} x^{-3/2} e^{-1/2x}, \quad x > 0 \quad \checkmark$$

From part (c), if $\tilde{N} \sim N(0, 1)$, then

$$\tilde{N}^2 \sim \chi\text{-square} \sim \Gamma\left(\frac{1}{2}, \frac{1}{2}\right).$$

Thus,

$$\tilde{X} = \tilde{N}^{-2} \sim \Gamma^{-1}\left(\frac{1}{2}, \frac{1}{2}\right) \quad \checkmark$$

To calculate the CDF we note that $\tilde{Y} = 1/\tilde{X}$ is a chi-squared random variable, whose CDF was calculated in part (c). We see that

$$F_{\tilde{X}}(x) = P(\tilde{X} < x)$$

$$= P\left(\frac{1}{\tilde{Y}} < x\right)$$

$$= P\left(\tilde{Y} > \frac{1}{x}\right)$$

$$= 1 - P\left(\tilde{Y} < \frac{1}{x}\right)$$

$$= 1 - F_{\tilde{Y}}\left(\frac{1}{x}\right)$$

$$= 1 - \text{erf}\left(\sqrt{\frac{1/x}{2}}\right)$$

$$= 1 - \text{erf}\left(\sqrt{\frac{1}{2x}}\right) = \text{erfc}\left(\sqrt{\frac{1}{2x}}\right) \quad \checkmark$$

by the definition of the complementary error function.

Problem 3. (a) For any $\Psi_1, \Psi_2 \in \mathcal{H}$,

$$\begin{aligned}\langle \Psi_1, \hat{P}_\Phi \Psi_2 \rangle &= \langle \Psi_1, \langle \Phi, \Psi_2 \rangle \Phi \rangle \\ &= \langle \Psi_1, \Phi \rangle \langle \Phi, \Psi_2 \rangle \\ &= \overline{\langle \Phi, \Psi_1 \rangle} \langle \Phi, \Psi_2 \rangle \\ &= \langle \langle \Phi, \Psi_1 \rangle \Phi, \Psi_2 \rangle \\ &= \langle \hat{P}_\Phi \Psi_1, \Psi_2 \rangle\end{aligned}$$

which is equivalent to

$$\hat{P}_\Phi^* = \hat{P}_\Phi. \quad \checkmark$$

On the other hand, for all $\Psi \in \mathcal{H}$

$$\begin{aligned}\hat{P}_\Phi^2 \Psi &= \hat{P}_\Phi (\hat{P}_\Phi \Psi) \\ &= \hat{P}_\Phi (\langle \Phi, \Psi \rangle \Phi) \\ &= \langle \Phi, \langle \Phi, \Psi \rangle \Phi \rangle \Phi \\ &= \langle \Phi, \Psi \rangle \langle \Phi, \Phi \rangle \Phi \\ &\quad \text{" } \|\Phi\|^2 = 1 \text{ ! } \\ &= \langle \Phi, \Psi \rangle \Phi = \hat{P}_\Phi \Psi\end{aligned}$$

$$\implies \hat{P}_\Phi^2 = \hat{P}_\Phi \quad \checkmark$$

(b) Since $\hat{I}^* = \hat{I}$ (easily), we see that

$$\begin{aligned}\hat{Q}^* &= (\hat{I} - \hat{P})^* \\ &= \hat{I}^* - \hat{P}^* = \hat{I} - \hat{P} = \hat{Q} \quad \checkmark\end{aligned}$$

Also,

$$\begin{aligned}\hat{Q}^2 &= (\hat{I} - \hat{P})^2 = \hat{I}^2 - 2\hat{I}\hat{P} + \hat{P}^2 \\ &= \hat{I} - 2\hat{P} + \hat{P} \quad (\text{using } \hat{P}^2 = \hat{P}) \\ &= \hat{I} - \hat{P} = \hat{Q} \quad \checkmark\end{aligned}$$

Furthermore,

$$\begin{aligned}\hat{P}\hat{Q} &= \hat{P}(\hat{I} - \hat{P}) \\ &= \hat{P} - \hat{P}^2 = \hat{P} - \hat{P} = 0 \quad \checkmark\end{aligned}$$

and, likewise,

$$\begin{aligned}\hat{Q}\hat{P} &= (\hat{I} - \hat{P})\hat{P} \\ &= \hat{P} - \hat{P}^2 = \hat{P} - \hat{P} = 0 \quad \checkmark\end{aligned}$$

NOTE: This last property implies that, for any $\Psi_1, \Psi_2 \in \mathcal{H}$,

$$\begin{aligned}\langle \hat{Q}\Psi_1, \hat{P}\Psi_2 \rangle &= \langle \Psi_1, \hat{Q}^* \hat{P} \Psi_2 \rangle \\ &= \langle \Psi_1, \hat{Q} \hat{P} \Psi_2 \rangle = 0\end{aligned}$$

so that $\hat{Q}\Psi_1 \perp \hat{P}\Psi_2$. This explains the notation $\hat{Q} = \hat{P}^\perp$.

(c) From the definition of \hat{Q} ,

$$\hat{I} = \hat{P} + \hat{Q}.$$

Thus,

$$\Psi = \hat{I}\Psi = (\hat{P} + \hat{Q})\Psi = \hat{P}\Psi + \hat{Q}\Psi$$

Since $\mathcal{A} = \text{range}(\hat{P})$, clearly $\hat{P}\Psi \in \mathcal{A}$!

Also, $\hat{P}\Psi$ is an eigenvector of \hat{P} with eigenvalue $= 1$,

$$\text{since } \hat{P}(\hat{P}\Psi) = \hat{P}^2\Psi = \hat{P}\Psi = +1 \cdot (\hat{P}\Psi).$$

Next we note that any $\Phi \in \mathcal{A}$ can be written, by definition as $\Phi = \hat{P}\Phi_0$ for some $\Phi_0 \in \mathcal{H}$. Then, $\forall \Phi \in \mathcal{A}$.

$$\begin{aligned} \langle \hat{Q}\Psi, \Phi \rangle &= \langle \hat{Q}\Psi, \hat{P}\Phi_0 \rangle \\ &= \langle \Psi, \hat{Q}^* \hat{P}\Phi_0 \rangle \\ &= \langle \Psi, \hat{Q}\hat{P}\Phi_0 \rangle = 0 \end{aligned}$$

Thus, $\hat{Q}\Psi \in \mathcal{A}^\perp$, by definition. ✓

$$\text{Finally, } \|\Psi\|^2 = \langle \Psi, \Psi \rangle$$

$$= \langle \hat{P}\Psi + \hat{Q}\Psi, \hat{P}\Psi + \hat{Q}\Psi \rangle$$

$$= \langle \hat{P}\Psi, \hat{P}\Psi \rangle + \langle \hat{Q}\Psi, \hat{P}\Psi \rangle = 0$$

$$+ \langle \hat{P}\Psi, \hat{Q}\Psi \rangle + \langle \hat{Q}\Psi, \hat{Q}\Psi \rangle$$

$$= \|\hat{P}\Psi\|^2 + \|\hat{Q}\Psi\|^2. \quad \checkmark$$

(d) By definition

$$\begin{aligned}\mathcal{A} &= \text{range}(\hat{P}) \\ &= \{ \hat{P}\Psi \mid \Psi \in \mathcal{H} \}\end{aligned}$$

For the projection in part (a),

$$\hat{P}\Psi = \langle \Phi, \Psi \rangle \Phi = c \Phi$$

with $c = \langle \Phi, \Psi \rangle \in \mathbb{C}$. Thus, every element of \mathcal{A} has the required form. To complete the proof, we must only show that any value $c \in \mathbb{C}$ is allowed.

But note that, for any choice of c ,

$$\begin{aligned}\hat{P}(c\Phi) &= c \hat{P}\Phi \\ &= c \underbrace{\langle \Phi, \Phi \rangle}_{=1} \Phi \\ &= c \Phi\end{aligned}$$

and thus

$$c\Phi \in \mathcal{A}.$$

We have thus shown that $\Psi \in \mathcal{A} = \text{range}(\hat{P})$ if and only if $\Psi = c\Phi$ for some $c \in \mathbb{C}$.